

Probability theory

Lesson 22

The Poisson Distribution

22.1- The Poisson Time Distribution.

22.1 - Problem 1:

Step 1: $\mu_t = \mu t = 3$, the average of 3 students at a local university violate curfew restrictions where $t = 2$.

Step 2: $k = 1$, the number of students that violate the curfew restrictions.

$$\text{Step 3: } P\{X = 1\} = \frac{(\mu t)^k}{k!} e^{-\mu t} \approx \frac{3^1}{1!} 2.718^{-3} \approx 0.15.$$

22.1 - Problem 2:

Step 1: $k = 5$, since we are interested in the probability that exactly 5 customers request an oil change.

Step 2: The time period we are concerned with is $t = 3$ hours. Therefore, we need to compute μ .

From the statement of the problem, we know that $2\mu = 3$. Therefore,

$\mu = 3/2$, average number of oil change requests per hour.

Since $t = 3$, we have $t\mu = 3(1.5) = 4.5$, average number of oil change requests over 3 hours.

$$\text{Step 3: } P\{X = 5\} = \frac{(\mu t)^k}{k!} e^{-\mu t} \approx \frac{(4.5)^5}{5!} 2.718^{-4.5} \approx 0.17 \text{ the probability that 5 customers request an}$$

oil change over 3 hours.

22.1 - Problem 3:

►(a).

Step 1: Compute μ .

Since $\mu_t = 12\mu = 6$, then $\mu = 0.5$ average number of fires per month.

Step 2: Compute the average number of fires over 2 months gives

$\mu_t = \mu t = 2\mu = 2(0.5) = 1$, the average number of fires over 2 months.

Step 3: To compute the probability that at least 1 fire occurs over 2 months, we use the formula

$$P\{X \geq 1\} = 1 - P\{X \leq 0\} = 1 - P\{X = 0\}.$$

Therefore, we need to use the Poisson formula for $k = 0$.

For $k = 0$, and $t = 2$ months, we have

$$P\{X = 0\} = \frac{(\mu t)^k}{k!} e^{-\mu t} \approx \frac{(1)^0}{0!} 2.718^{-1} \approx 0.362.$$

►(b).

The probability that at most 1 fire can occur over 2 months can be written

$$P\{X \leq 1\} = P\{X = 0\} + P\{X = 1\}.$$

From (a), we know that

$$\mu_t = \mu t = 2\mu = 2(0.5) = 1,$$

the average number of fires over 2 months.

Since we need to compute

$$P\{X = 0\}$$

and

$$P\{X = 1\},$$

we have

$$P\{X = 0\} = \frac{(\mu t)^k}{k!} e^{-\mu t} \approx \frac{(1)^0}{0!} 2.718^{-1} \approx 0.36$$

$$P\{X = 1\} = \frac{(\mu t)^k}{k!} e^{-\mu t} \approx \frac{(1)^1}{1!} 2.718^{-1} \approx 0.36$$

Therefore,

$$P\{X \leq 1\} = P\{X = 0\} + P\{X = 1\} \approx 0.36 + 0.36 = 0.72.$$

22.1 - Problem 4:

Step 1: The average number of orders returned over $t = 7$ days is

$$\mu_t = \mu t = (3\%)2000 = (0.03)2000 = 60.$$

Step 3: Computing the average number of orders returned over 1 day gives

$$\mu_t = \mu t = 60$$

$$t = 7$$

$$\mu_t = \mu t = \mu 7 = 60$$

$\mu = 60/7$, average number of orders returned over 1 day.

Here $t = 1$.

Step 5: The number is orders returned is $k = 5$.

$$\text{Step 6: } P\{X = 5\} = \frac{(\mu t)^k}{k!} e^{-\mu t} \approx \frac{\left(\frac{60}{7}\right)^5}{5!} 2.718^{-\frac{60}{7}} \approx 0.073.$$

22.2 - The Poisson Spatial Distribution.

22.1 - Problem 1:

Step 1: Let $t = 10$, the number of pages.

Step 2: $\mu = 0.25$

Step 3: The average number of errors over 10 pages is

$$\mu_t = \mu t = 0.25(10) = 2.5.$$

Step 4: The number of errors is

$$k = 2.$$

Step 5: The probability that only 2 errors occur over 10 pages is

$$P\{X = k\} = \frac{(\mu t)^k}{k!} e^{-\mu t} = P\{X = 2\} \approx \frac{(2.5)^2}{2!} 2.718^{-2.5} \approx 0.26.$$

22.2 - Problem 2:

Step 1: The average number of coaxial cable found defective, for every 100,000 feet stalled
 $100,000\mu = 10$.

Step 2: Computing μ , $\mu = \frac{10}{100000} = 0.0001$, average per foot.

Step 3: The average number of coaxial cable found defective, for every 10,000 feet stalled is

$$t = 10,000$$

$$\mu_t = (10,000)0.0001 = 1.$$

Step 4: The probability that 2 feet are defective is

$$P\{X = k\} = \frac{(\mu t)^k}{k!} e^{-\mu t} = P\{X = 2\} \approx \frac{1^2}{2!} 2.718^{-2} \approx 0.07.$$

22.2 - Problem 3:

►(a).

Step 1: The average number of clients that like the computer is

$$\mu_t = \mu t = 35\%(1,000) = (0.35)(1000) = 350 \text{ and}$$

$$\mu = 0.35$$

Step 2: The average number of clients out of the 10 sent that like the computer is

$$\mu_t = 10\mu = 10(0.35) = 3.5.$$

Step 3: The probability of not continuing sales on the computer is

$$P\{X \leq 1\} = P\{X = 0\} + P\{X = 1\} \approx \frac{3.5^0}{0!} 2.718^{-3.5} + \frac{3.5^1}{1!} 2.718^{-3.5} \approx 0.136$$

►(b).

Step 1: The average number of clients that like the computer is

$$\mu_t = \mu t = 20\%(1,000) = (0.20)(1000) = 200 \text{ and}$$

$$\mu = 0.20$$

Step 2: The average number of clients out of the 10 sent that like the computer is

$$\mu_t = 10\mu = 10(0.20) = 2$$

Step 3: The probability of continuing sales on the computer is

$$P\{X \geq 2\} = 1 - P\{X = 0\} - P\{X = 1\} \approx 1 - \frac{2^0}{0!}2.718^{-2} - \frac{2^1}{1!}2.718^{-2} \approx 0.594.$$

Supplementary Problems

1.

Step 1: $t = 100$, and $\mu_t = t\mu = 3$, average per box.

$$\text{Step 2: } P\{X = k\} = p = \frac{(\mu t)^k}{k!} e^{-\mu t} = P\{X = 1\} \approx \frac{3^1}{1!} 2.718^{-3} \approx 0.149,$$

the probability that a single box contains exactly one broken bulb.

Step 3: Applying the binomial distribution for

$$N = 5,$$

$$k = 3$$

$$p = 0.149$$

we have

$$P\{Y = 3\} = \binom{5}{3} (0.149)^3 (0.851)^2 \approx 10(0.0033)(0.724) \approx 0.024.$$

2.

► a.

Step 1: **E**: The event that the box contains 1 defective transmission.

W: The event that the transmission came from the West coast.

W': The event that the transmission came from the East coast.

X: The random variable of the number of defective parts in a box.

$$P(\mathbf{W}') = \frac{10,000}{25,000} = 0.40$$

$$P(\mathbf{W}) = 0.60$$

$$\text{Step 2: } \mathbf{E} = (\mathbf{E} \cap \mathbf{W}) \cup (\mathbf{E} \cap \mathbf{W}')$$

$$P(\mathbf{E}) = P(\mathbf{E} \cap \mathbf{W}) + P(\mathbf{E} \cap \mathbf{W}')$$

$$P(\mathbf{E} \cap \mathbf{W}) = P(\mathbf{E} | \mathbf{W})P(\mathbf{W}) = P(\mathbf{X} = 1 | \mathbf{W})P(\mathbf{W}) = P(\mathbf{X} = 1 | \mathbf{W})(0.60)$$

$$P(\mathbf{E} \cap \mathbf{W}') = P(\mathbf{E} | \mathbf{W}')P(\mathbf{W}') = P(\mathbf{X} = 1 | \mathbf{W}')P(\mathbf{W}') = P(\mathbf{X} = 1 | \mathbf{W}')(0.40)$$

Step 3: For the East coast the average number of defective parts is

$$\mu = 500/10,000 = 0.05.$$

Therefore, the average number of defective transmission per box from the East coast is

$$\mu_t = \mu t = 0.05(100) = 5$$

For 1 defective part we have $k = 1$ and therefore,

$$P\{X = k | \mathbf{W}'\} = \frac{(\mu t)^k}{k!} e^{-\mu t} = P\{X = 1 | \mathbf{W}'\} \approx \frac{5^1}{1!} 2.718^{-5} \approx 0.03,$$

the probability that the box contains one defective transmission.

Step 4: For the West coast the average number of defective parts is

$$\mu = 450/15,000 = 0.03.$$

Therefore, the average number of defective transmission per box from the West coast is

$$\mu_t = \mu t = 0.03(100) = 3.$$

For 1 defective part, we have $k = 1$ and therefore

$$P\{X = k | \mathbf{W}\} = \frac{(\mu t)^k}{k!} e^{-\mu t} = P\{X = 1 | \mathbf{W}\} \approx \frac{3^1}{1!} 2.718^{-3} \approx 0.15,$$

the probability that the box contains exactly 1 defective transmission.

$$\text{Step 5: } P(\mathbf{E}) = P(\mathbf{E} \cap \mathbf{W}) + P(\mathbf{E} \cap \mathbf{W}') = P(\mathbf{X} = 1 | \mathbf{W})(0.60) + P(\mathbf{X} = 1 | \mathbf{W}')(0.40) =$$

$$(0.15)(0.60) + (0.03)(0.40) = 0$$

$$\mathbf{D} = (\mathbf{D} \cap \mathbf{E}) \cup (\mathbf{D} \cap \mathbf{W})$$

$$P(\mathbf{D}) = P(\mathbf{D} \cap \mathbf{E}) + P(\mathbf{D} \cap \mathbf{W}) = P(\mathbf{E})P(\mathbf{D}|\mathbf{E}) + P(\mathbf{W})P(\mathbf{D}|\mathbf{W}) = (0.40)(0.03) + (0.60)(0.15) = 0.102$$

►b.

This is a Bayesian problem:

$$P(\mathbf{W}'|\mathbf{E}) = \frac{P(\mathbf{E}|\mathbf{W}')P(\mathbf{W}')}{P(\mathbf{E})} \approx \frac{(0.03)(0.4)}{0.10} = 0.12$$

3.

Using the Poisson distribution ,

$$P\{X = k\} = \frac{(\mu t)^k}{k!} e^{-\mu t},$$

we compute the following table:

k	$P\{X = k\} = \frac{(\mu t)^k}{k!} e^{-\mu t}$
0	0.01
1	0.045
2	0.10
3	0.15
4	0.17
5	0.15

Therefore, $k = 4$.

4.

►a.

Step 1: $t = 1$ page

$\mu = 1$ per page ,

$\mu_t = t\mu = 1$

Step 2: $P\{X \geq 3\} = 1 - P\{X \leq 2\} = 1 - [P\{X=0\} + P\{X=1\} + P\{X=2\}]$

$$P\{X = k\} = \frac{(\mu t)^k}{k!} e^{-\mu t} = \frac{e^{-1}}{k!}$$

$$P\{X = 0\} = \frac{e^{-1}}{0!} = e^{-1} \approx 0.37$$

$$P\{X = 1\} = \frac{e^{-1}}{1!} = e^{-1} \approx 0.37$$

$$P\{X = 2\} = \frac{e^{-1}}{2!} = 0.37/2 \approx 0.185$$

$$\text{Step 3: } P\{X \geq 3\} \geq 1 - 0.37 - 0.37 - 0.185 \approx 0.08$$

►b.

Step 1:

$t = 3$ pages

$\mu = 1$ per page ,

$\mu_t = t\mu = 3$

$$P\{X = 0\} = \frac{(\mu t)^k}{k!} e^{-\mu t} = \frac{3^0 e^{-3}}{0!} = e^{-3} \approx 0.05$$

5.

The basic unit is 1 person.

$\mu = 0.01$, average per person.

$t = N$, the sample size

$\mu_t = \mu t = (0.01)N$

$$P\{X \geq 1\} = 1 - P\{X = 0\}$$

$$P\{X = 0\} = \frac{(0.01N)^0}{0!} e^{-0.01N} = e^{-0.01N}$$

$$P\{X \geq 1\} = 1 - e^{-0.01N} \geq 0.95$$

$$1 - 0.95 \geq e^{-0.01N}$$

$$0.05 \geq e^{-0.01N}$$

$$\ln(0.05) \geq \ln(e^{-0.01N}) = -0.01N$$

$$\ln(.05) \approx -3$$

$$-3 \geq -0.01N$$

$$300 \leq N$$

6.

X: The random variable that equals the number of sales.

$$P\{X \geq 1\} \geq 0.99$$

$$P\{X \geq 1\} = 1 - P\{X = 0\} \geq 0.99$$

$$1 - 0.99 \geq P\{X = 0\}$$

$$0.01 \geq P\{X = 0\} = \frac{(\mu t)^0}{0!} e^{-\mu t} = e^{-\mu t}$$

$$e^{\mu t} \geq 1/0.01 = 100$$

$$t = 1$$

$$e^{\mu} \geq 100$$

$$\ln(e^{\mu}) \geq \ln(100) \approx 5$$

$$\ln(e^{\mu}) = \mu \approx 5$$

7.

► a.

Step 1: First we need to find the probability p of a hand having a full house.

$$\#S = \binom{52}{5} = 2,598,960$$

E: A full house occurs.

$$\#E = 13 \binom{4}{3} 12 \binom{4}{2} = 3744$$

$$P(E) = p = \frac{3744}{2598960} = \frac{6}{4165}$$

Step 2: $t = 10$

$$\mu = p = \frac{6}{4165}, \text{ per hand}$$

$$\mu t = 10 \left(\frac{6}{4165} \right) = \frac{12}{833}$$

$$P\{X \geq 1\} = 1 - P\{X = 0\} = 1 - \frac{\left(\frac{12}{833}\right)^0}{0!} e^{-\frac{12}{833}} = 1 - e^{-\frac{12}{833}} \approx 1 - 0.99 = 0.01$$

► b.

N: Number of hands

$$\mu N = \left(\frac{6}{4165} \right) N$$

$$P\{X \geq 1\} > 1/2$$

$$P\{X = 0\} = \frac{\left(\frac{6}{4165}N\right)^0}{0!} e^{-\frac{6}{4165}N} = e^{-\frac{6}{4165}N}$$

$$P\{X \geq 1\} = 1 - P\{X = 0\} = 1 - e^{-\frac{6}{4165}N} > 1/2$$

$$e^{-\frac{6}{4165}N} < 1/2$$

$$2 < e^{\frac{6}{4165}N}$$

$$\ln(2) < N(6/4165)$$

$$(0.69)(4165/6) < N$$

$$478 < N$$

8.

►a.

We need to use the formula, from supplementary 16, lesson 17 where $N = 5$,

$$\binom{N}{2} ({}_{365}P_{N-1}) 365^{-N}$$

is the probability that for N random people, 2 are born on the same day and the others on different days.

For our problem, let $N = 10$. Therefore, the probability that for a given group,

$$p = \binom{10}{2} ({}_{365}P_9) 365^{-10} = 45 {}_{365}P_9 365^{-10} \approx 45(0.002) = 0.09.$$

Now we apply the Poisson distribution:

$$t = 100$$

$$\mu = 0.09$$

$$\mu t = 9$$

$$k = 10\%(100) = 0.10(100) = 10$$

$$P\{X = 10\} = \frac{(\mu t)^k}{k!} e^{-\mu t} = \frac{9^{10} e^{-9}}{10!} \approx 0.12.$$

►b.

We need to use the formula, from supplementary 16, lesson 17 where $N = 5$,

$$\binom{N}{2} ({}_{365}P_{N-1}) 365^{-N}$$

is the probability that for N random people, 2 are born on the same day and the others on different days.

For our problem, let $N = 5$. Therefore, the probability that for a given group,

$$p = \binom{5}{2} ({}_{365}P_4) 365^{-5} = 10 {}_{365}P_4 365^{-5} \approx 10(0.003) = 0.03.$$

Now we apply the Poisson distribution:

$$t = 200$$

$$\mu = 0.03$$

$$\mu t = 6$$

$$k = 10\%(200) = 0.10(200) = 20$$

$$P\{X = 20\} = \frac{(\mu t)^k}{k!} e^{-\mu t} = \frac{6^{20} e^{-6}}{20!} \approx 0.$$

9.

Step 1:

E: The event that at least 1 page of the book contains more than k errors.

E': The event that each page of the book contains at most k errors.

A_j: The event that the j th page of the book contains at most k errors, where each of these pages are independent.

Step 2: Let X be a random variable with a Poisson distribution such that

$$P(X = v) = \frac{(\lambda t)^v}{v!} e^{-\lambda t}.$$

Since $t = 1$ represents v errors per page, we have

$$P(X = v) = \frac{(\lambda)^v}{v!} e^{-\lambda} = p(v, \lambda); v = 0, 1, 2, \dots$$

Therefore,

$$P(\mathbf{A}_j) = P(X \leq k) = P(X = 0) + P(X = 1) + \dots + P(X = k) = p(0, \lambda) + p(1, \lambda) + \dots + p(k, \lambda) = p.$$

$$E' = A_1 \cap A_2 \cap A_3 \cap \dots \cap A_n$$

From independence, $P(E') = P(A_1 \cap A_2 \cap A_3 \cap \dots \cap A_n) = P(A_1)P(A_2)P(A_3)\dots P(A_n) = p^n$,

$$P(E) = 1 - p^n.$$

10.

► a.

To compute μ , we complete the table below:

k: Number of false alarms	0	1	2	3	4	5
H_k: Number of hours where k false alarms occurred	29	27	11	4	1	0
kH_k:	0	27	22	12	4	0

$$\mu = (0 + 27 + 22 + 12 + 4 + 0)/72 \approx 0.903$$

► b.

k: Number of false alarms	0	1	2	3	4	5
$H_k = (72) \frac{(0.903)^k}{k!} e^{-0.903}$	29.18	26.35	11.90	3.58	1	0
Number of hours where k false alarms occurred						

11.

► a.

To compute μ , we complete the table below:

k: Number of infections	0	1	2	3	4	6	7	8	9
A_k: Number of areas where k infections exists.	200	112	80	51	27	15	4	6	5
kA_k:	0	112	160	153	108	90	28	48	45

$$\mu = (0 + 112 + 160 + 153 + 108 + 90 + 28 + 48 + 45)/500 \approx 1.49$$

►b.

k: Number of infections	0	1	2	3	4	5	6	7	8	9
$A_k =$ $(500) \frac{(1.49)^k}{k!} e^{-1.49}$	112.69	167.9	125.09	62.13	23.14	6.9	1.71	0	0	0
Number of areas where k infections exists										

12.

$$P(X = k) = [(\mu t)^k e^{-\mu t}] / k!$$

For this problem $\mu = 3$, $t = 2$, per 100,000 residents.

Therefore,

$$P(X = k) = [(6)^k e^{-6}] / k!$$

►a.

$$k = 8$$

$$P(X = 8) = [(6)^8 e^{-6}] / 8! \approx 0.1033$$

►b.

$$P(4 \leq X \leq 8) = P(X = 4) + P(X = 5) + P(X = 6) + P(X = 7) + P(X = 8) =$$

$$[(6)^4 e^{-6}] / 4! + [(6)^5 e^{-6}] / 5! + [(6)^6 e^{-6}] / 6! + [(6)^7 e^{-6}] / 7! + [(6)^8 e^{-6}] / 8! \approx 0.696$$

►c.

$$P(X < 3) = P(X = 0) + P(X = 1) + P(X = 2) = [(6)^0 e^{-6}] / 0! + [(6)^1 e^{-6}] / 1! + [(6)^2 e^{-6}] / 2! \approx 0.062$$

►d.

$$P(X > 3) = 1 - [P(X = 0) + P(X = 1) + P(X = 2) + P(X = 3)] =$$

$$1 - [(6)^0 e^{-6}] / 0! + [(6)^1 e^{-6}] / 1! + [(6)^2 e^{-6}] / 2! + [(6)^3 e^{-6}] / 3! \approx 1 - 0.151 = 0.849$$

13.

$$P(X = k) = [(\mu t)^k e^{-\mu t}] / k!$$

For this problem $\mu = 2.5$, $t = 1$, per minute. Therefore,

$$P(X = k) = [(2.5)^k e^{-2.5}] / k! .$$

►a.

$$P(X = 3) = [(2.5)^3 e^{-2.5}] / 3! \approx 0.2138$$

►b.

$$P(X < 5) = P(X = 0) + P(X = 1) + P(X = 2) + P(X = 3) + P(X = 4) =$$

$$(2.5)^0 e^{-2.5} / 0! + (2.5)^1 e^{-2.5} / 1! + (2.5)^2 e^{-2.5} / 2! + (2.5)^3 e^{-2.5} / 3! + (2.5)^4 e^{-2.5} / 4! \approx 0.8912$$

►c.

$$P(X > 6) = 1 - P(X \leq 6) =$$

$$1 - \{P(X = 0) + P(X = 1) + P(X = 2) + P(X = 3) + P(X = 4) + P(X = 5) + P(X = 6)\} =$$

$$1 - \{(2.5)^0 e^{-2.5} / 0! + (2.5)^1 e^{-2.5} / 1! + (2.5)^2 e^{-2.5} / 2! + (2.5)^3 e^{-2.5} / 3! + (2.5)^4 e^{-2.5} / 4! + (2.5)^5 e^{-2.5} / 5! + (2.5)^6 e^{-2.5} / 6!\} \approx$$

$$0.0142$$

14.

►a.

$$P(X \geq 2 | X \leq 4) = P[(X \leq 4) \cap (X \geq 2)] / P(X \leq 4) = P(2 \leq X \leq 4) / P(X \leq 4)$$

$$P(2 \leq X \leq 4) = P(X = 2) + P(X = 3) + P(X = 4) = 2^2 e^{-2} / 2! + 2^3 e^{-2} / 3! + 2^4 e^{-2} / 4! \approx 0.271 + 0.180 + 0.09$$

$$= 0.541$$

$$P(0 \leq X \leq 4) = P(X = 0) + P(X = 1) + P(X = 2) + P(X = 3) + P(X = 4) \approx$$

$$0.135 + 0.271 + 0.271 + 0.180 + 0.09 \approx 0.947$$

$$P(X \geq 2 | X \leq 4) = 0.541 / 0.947 \approx 0.571$$

►b.

$$P(X \leq 4 | X \geq 2) = P[(X \leq 4) \cap (X \geq 2)] / P(X \geq 2) = P(2 \leq X \leq 4) / P(X \geq 2)$$

$$P(X \geq 2) = 1 - P(X \leq 1) \approx 1 - 0.406 = 0.594$$

15.

►a. We will use the binomial expansion: $(\lambda + \mu)^m =$

$$\binom{m}{0} \lambda^m \mu^0 + \binom{m}{1} \lambda^{m-1} \mu^1 + \binom{m}{r} \lambda^{m-r} \mu^r + \dots + \binom{m}{m} \lambda^0 \mu^m$$

where $\binom{m}{r} = \frac{m!}{r!(m-r)!}$

$Z = X + Y$

$P(Z=m) = P(X=m)P(Y=0) + P(X=m-1)P(Y=1) + \dots + P(X=m-r)P(Y=r) + \dots + P(X=0)P(Y=m)$

$$= \frac{\lambda^m \mu^0}{m!0!} e^{-\lambda} e^{-\mu} + \frac{\lambda^{m-1} \mu^1}{(m-1)!1!} e^{-\lambda} e^{-\mu} + \dots + \frac{\lambda^{m-r} \mu^r}{(m-r)!r!} e^{-\lambda} e^{-\mu} + \dots + \frac{\lambda^0 \mu^m}{(0)!m!} e^{-\lambda} e^{-\mu} =$$

$$e^{-\lambda} e^{-\mu} \left[\frac{\lambda^m \mu^0}{m!0!} + \frac{\lambda^{m-1} \mu^1}{(m-1)!1!} + \dots + \frac{\lambda^{m-r} \mu^r}{(m-r)!r!} + \dots + \frac{\lambda^0 \mu^m}{(0)!m!} \right] =$$

$$\frac{e^{-(\lambda + \mu)}}{m!} \left[\frac{m! \lambda^m \mu^0}{m!0!} + \frac{m! \lambda^{m-1} \mu^1}{(m-1)!1!} + \dots + \frac{m! \lambda^{m-r} \mu^r}{(m-r)!r!} + \dots + \frac{m! \lambda^0 \mu^m}{(0)!m!} \right] =$$

$$\frac{e^{-(\lambda + \mu)}}{m!} \left[\binom{m}{0} \lambda^m \mu^0 + \binom{m}{1} \lambda^{m-1} \mu^1 + \dots + \binom{m}{r} \lambda^{m-r} \mu^r + \dots + \binom{m}{m} \lambda^0 \mu^m \right] =$$

$$\frac{e^{-(\lambda + \mu)}}{m!} (\lambda + \mu)^m = P(Z = m)$$

►b. $P(X = k | X + Y = m) = P[(X = k) \cap (X + Y = m)] / P(X + Y = m)$

$P[(X = k) \cap (X + Y = m)] = P[(X = k) \cap (Y = m - k)] = P(X = k)P(Y = m - k) =$

$$\frac{e^{-\lambda} \lambda^k}{k!} \frac{e^{-\mu} \mu^{m-k}}{(m-k)!}$$

$P(X + Y = m) = \frac{e^{-(\lambda + \mu)}}{m!} (\lambda + \mu)^m$

$$P(X = k | X + Y = m) = \frac{\frac{e^{-(\lambda + \mu)}}{k!(m-k)!} \lambda^k \mu^{m-k}}{\frac{e^{-(\lambda + \mu)} (\lambda + \mu)^m}{m!}} = \frac{m! \lambda^k \mu^{m-k}}{k!(m-k)! (\lambda + \mu)^m} = \frac{\binom{m}{k} \lambda^k \mu^{m-k}}{(\lambda + \mu)^m} =$$

$$\frac{\binom{m}{k} \lambda^k \mu^{m-k}}{(\lambda + \mu)^k (\lambda + \mu)^{m-k}} = \binom{m}{k} \left(\frac{\lambda}{\lambda + \mu}\right)^k \left(\frac{\mu}{\lambda + \mu}\right)^{m-k}$$

16.

► a.

The Poisson distribution for robot A is $P(X = k) = \frac{e^{-0.23}}{k!} (0.23)^k$

The Poisson distribution for robot B is $P(Y = k) = \frac{e^{-0.28}}{k!} (0.28)^k$

From Problem 15a, we have the Poisson distribution for the total number of defects detected by both robots per automobile is

$$P(X + Y = m) = \frac{e^{-0.23 + 0.28}}{m!} (0.23 + 0.28)^m = \frac{e^{-0.51}}{m!} (0.51)^m$$

Step 1: First we complete row 2 to 3 places of accuracy :

Number of k defects	0	1	2	3	4	5
P(X + Y) = k	0.6	0.306	0.078	0.013	0.002	0
Number of autos with k defects						

Step 2: The total number of automobiles assembled is 575.

To compute the third row, we multiply 575 by the probabilities in the second row:

Number of k defects	0	1	2	3	4	5
P(X + Y) = k	0.6	0.306	0.078	0.013	0.002	0
Number of autos with k defects	345	176	45	8	1	0

► b.

From Problem 15b, we have

$$P(X = k | X + Y = m) = \binom{m}{k} \left(\frac{\lambda}{\lambda + \mu}\right)^k \left(\frac{\mu}{\lambda + \mu}\right)^{m-k}$$

where

$$\lambda = 0.23,$$

$$\mu = 0.28,$$

$$m = 3, k = 1$$

$$P(X = 1|X + Y = 3) = \binom{3}{1} \left(\frac{0.23}{0.51}\right)^1 \left(\frac{0.28}{0.51}\right)^2 \approx 0.41.$$

►c. By Bayes theorem we have:

$$P(X + Y = m| X = k) = P(X = k|X + Y = m)P(X + Y = m)/P(X = k) =$$

$$P(X + Y = 3| X = 1) = P(X = 1|X + Y = 3)P(X + Y = 3)/P(X = 1).$$

From above we have:

$$P(X = 1|X + Y = 3) \approx 0.41, P(X + Y = 3) = 0.013 \text{ and}$$

$$P(X = 1) = \frac{e^{-0.23}}{1!} (0.23)^1 \approx 0.18$$

Which gives us

$$P(X + Y = 3| X = 1) \approx (0.41)(0.013)/0.18 = 0.03.$$

The exponential distribution. A random variable X has a exponential distribution if

$$P(X \leq x) = 1 - e^{-\lambda x} \text{ for } x \geq 0 \text{ and}$$

$$P(X < x) = 0 \text{ for } x < 0.$$

17.

Show how the exponential distribution is related to the Poisson distribution. Also give an interpretation of the meaning of the exponential distribution.

The exponential distribution is

$$P(Y = k) = \frac{(\mu t)^k}{k!} e^{-\mu t}$$

Let $k = 0$.

This gives

$$P(Y = 0) = e^{-\mu t}$$

Let $\mu = \lambda$.

$$P(Y = 0) = e^{-\lambda t}$$

For $x \geq 0$, $P(X \leq x) = 1 - e^{-\lambda x} = 1 - P(Y = 0)$.

Assume $t = x$ where t is time. This gives us

$$\text{for } t \geq 0, P(X \leq t) = 1 - e^{-\lambda x} = 1 - P(Y = 0).$$

Assume we have an event that occurs under the Poisson distribution. The event $(Y = 0)$ means that up to the time t , the event has not yet occurred. Since

$$P(X \leq t) = 1 - e^{-\lambda x} = 1 - P(Y = 0),$$

$$(X \leq t)$$

means that by the time t occurs the event has occurred at least once.

Therefore, $P(X \leq t)$ is the probability that an event has occurred at least once.

18.

► a.

$$P(X > u + v | X > u) = P[(X > u + v) \cap (X > u)] / P(X > u) = P(X > u + v) / P(X > u)$$

$$P(X > u + v) = e^{-\lambda(u+v)}$$

$$P(X > u) = e^{-\lambda u}$$

$$P(X > u + v | X > u) = P(X > u + v) / P(X > u) = e^{-\lambda(u+v)} / e^{-\lambda u} = e^{-\lambda v} = P(X > v)$$

► b.

This means that the random variable X 's outcome is independent of the time that it starts.

This means we can write

$$P(X > u + v - u | X > u - u) = P(X > v | X > 0) = P(X > v)$$

$$\mathbf{19.} P(X > u + v | X > u) = P(X > u + v) / P(X > u) = P(X > v)$$

$$P(X > u + v) = P(X > u)P(X > v)$$

Let $P(X > 1) = a$

$$P(X > 2) = P(X > 1 + 1) = P(X > 1)P(X > 1) = a^2$$

$$P(X > 3) = P(X > 2 + 1) = P(X > 2)P(X > 1) = a^2a$$

Continuing by induction:

$$P(X > n) = P(X > n - 1 + 1) = P(X > n - 1)P(X > 1) = a^{n-1}a = a^n$$

$$P(X > 1) = P(X > n/n) = P(X > 1/n + 1/n + \dots + 1/n) = [P(X > 1/n)]^n = a$$

$$P(X > 1/n) = a^{1/n}$$

$$P(X > m/n) = P(X > 1/n + 1/n + \dots + 1/n) = [P(X > 1/n)]^m = (a^{1/n})^m = a^{m/n} = (1/a)^{-m/n}$$

$$\text{Let } e^\lambda = 1/a$$

$$\ln(e^\lambda) = \ln(1/a) = \lambda$$

Since $1/a > 1$, we have $\lambda > 0$.

$$\text{Therefore, } P(X > m/n) = (1/a)^{-m/n} = (e^\lambda)^{-m/n} = e^{-\lambda m/n}$$

Now since the function

$$y = e^{-\lambda x}$$

is continuous for a values of x , we can define

$$P(X > x) = e^{-\lambda x}$$

for all irrational numbers x .

$$\text{Finally } P(X \leq x) = 1 - P(X > x) = 1 - e^{-\lambda x}$$